Dimension Reduction in Dynamical Systems using Factor Models

Dimension reduction in physical and data sciences

Sayan Mukherjee

https://sayanmuk.github.io/

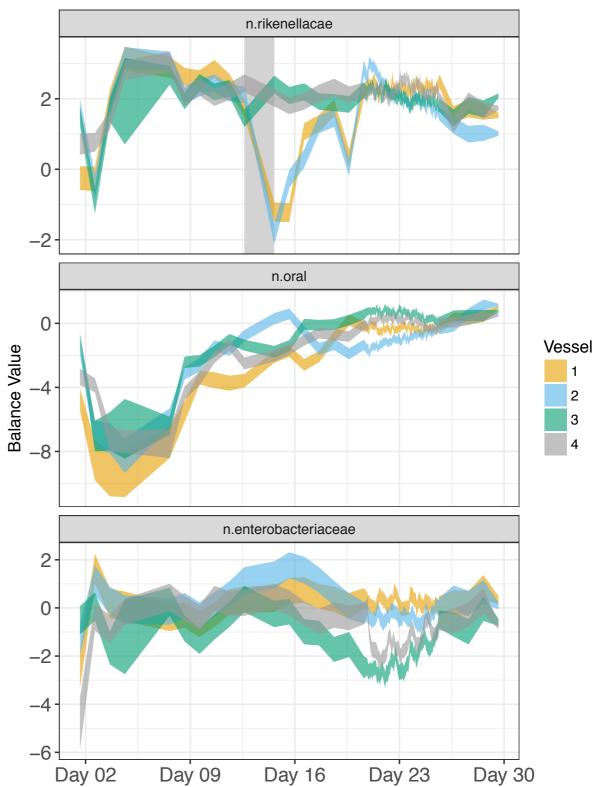
Joint work with:

— K. McGoff (UNC Ch) | D. Hoang (Duke) | A. Nobel (UNC CH)

Posterior Consistency for Dynamical Systems

Microbial ecology

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Posterior 95% credible interval

General framework

We consider mathematical models of the form:

- X is the "phase space";
- X_t is the "true state" of bioreactor (your stomach) at time t;
- Y is the "observation space";
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We only have access to the observations $\{Y_{t_k}\}_{k=0}^n$.

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- ▶ what is the "true state" of the bioreactor at time *t*? (filtering)
- what are we likely to observe at time t_{n+1} ? (prediction)
- what are the rules governing the evolution of the system? (model selection / parameter estimation)

We'll focus on the last type of question.

Basic assumptions

How are the variables $\{X_{t_k}\}_{k=0}^n$ and $\{Y_{t_k}\}_{k=0}^n$ related?

We'll assume the process $(X_t, Y_t)_t$ has:

stationarity: the rules governing both the state space and our observations don't change over time.

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- Markov property: given the microbial population today, the microbial population tomorrow is independent of the population yesterday.
- conditionally independent observations: given the state of the population today, today's observation is independent of any other variables.

Such systems are called "hidden Markov models" (HMMs).

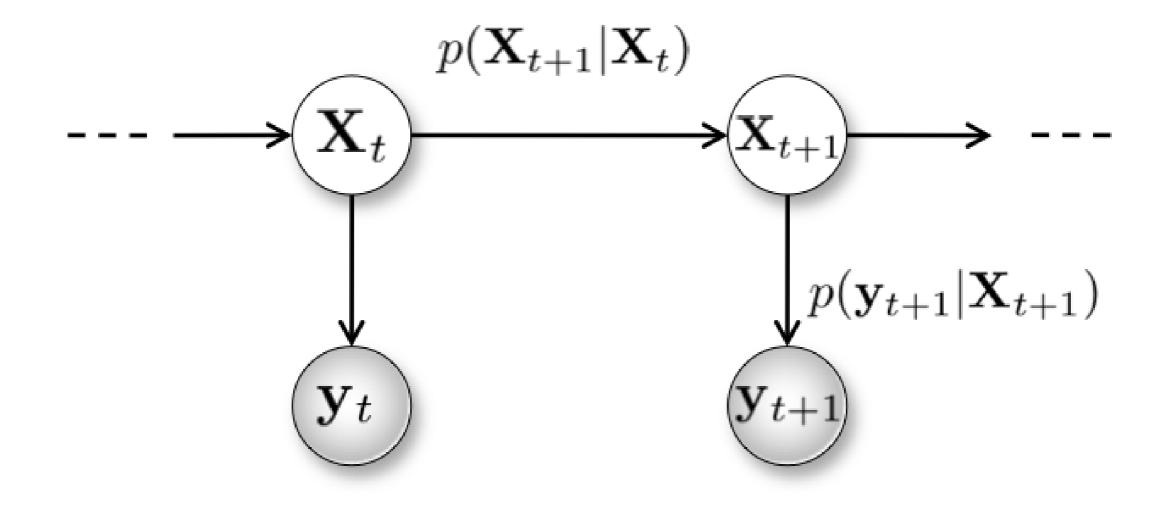
HMMs

The basic model for a HMM is

$$egin{array}{rcl} X_k &=& a_{ heta}(X_{k-1},W_k), \ Y_k &=& b_{ heta}(X_k,V_k), \end{array}$$

where $(V_k)_k$ and $(W_k)_k$ are sequences of iid random variables independent of X_0 .

HMMs



Stochastic versus deterministic systems

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- Otherwise, we'll say the process $(X_t)_t$ is deterministic.

In ecology both types of systems are commonly used.

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Family of systems $(X, X, T_{\theta}, \mu_{\theta})_{\theta \in \Theta} \equiv (T_{\theta}, \mu_{\theta})_{\theta \in \Theta}$.

Observational noise

Conditional likelihood: $g_{\theta}(y \mid x) = f(Y_t = y \mid x_t = x, \theta)$, with

$$\int g_{\theta}(y \mid x) d\nu(y) = 1.$$

Also $g: \Theta \times \mathsf{X} \times \mathsf{Y} \to \mathbb{R}_+$.

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Likelihood for y_0^n in Y^{n+1} conditioned on θ and $X_0 = x$ is

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Likelihood for y_0^n in Y^{n+1} conditioned on θ and $X_0 = x$ is

$$p_{ heta}(y_0^n \mid x) = \prod_{k=0}^n g_{ heta}(y_k \mid T_{ heta}^k(x)),$$

and the (marginal) likelihood of observing y_0^n given θ is

$$p_{\theta}(y_0^n) = \int p_{\theta}(y_0^n \mid x) d\mu_{\theta}(x).$$

An example

•
$$X_0 \sim U[0, 1];$$

• $X_{k+1} = \theta X_k (1 - X_k);$

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•
$$Y_k \sim N(X_k, \sigma_{\theta}^2)$$
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Approaches to estimation

There are many approaches to estimation:

- maximum likelihood estimation,
- Bayesian estimation,
- optimization (minimization of a cost function),
- ► etc.

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We'll focus on one approaches:

(2) Bayesian inference.

Preliminaries

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Tracking systems: Compact metrizable space $\mathcal{X} := X \times \Theta$ with map $S : \mathcal{X} \to \mathcal{X}$.

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$$S: \Theta \times X \to X, \quad S_{\theta}: X \to X.$$

Loss or regret: $\ell : \mathcal{X} \times \mathcal{Y} \to \mathbb{R}_+$. Cost of

$$\ell_n(x,y) := \ell_n(x_0^{n-1}, y_0^{n-1}) = \sum_{k=0}^{n-1} \ell(x_k, y_k),$$

 $x_0^{n-1} = (x, Sx, \dots, S^{n-1}x) \text{ and } y_0^{n-1} = (y, Ty, \dots, T^{n-1}y).$

Dynamic linear models

$$\begin{array}{rcl} x_{t+1} &=& A_{t+1} x_t \\ y_t &=& B_t x_t + v_t, \end{array}$$

Here:

 y_t is an observation in \mathbb{R}^p ; x_t is a hidden state in \mathbb{R}^q ;

 A_t is a $p \times p$ state transition matrix;

 B_t is a $q \times p$ observation matrix;

 v_t is a zero-mean vector in \mathbb{R}^q .

Likelihood: Lik(data | θ)

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$$\pi(\theta \mid \mathsf{data}) = \frac{\mathsf{Lik}(\mathsf{data} \mid \theta) \times \pi(\theta)}{\mathsf{Pr}(\mathsf{data})}$$

Gibbes:

$$\pi(heta \mid \mathsf{data}) = rac{\exp(-\ell(\mathsf{data} \mid heta)) imes \pi(heta)}{\int_{ heta} \exp(-\ell(\mathsf{data} \mid heta)) imes \pi(heta) d heta}.$$

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Consider the probability measure over Borel sets $A \subset \mathcal{X}$

$$P_n(A \mid y) = \frac{\int_A \exp(-\ell_n(x, y)) d\pi(x)}{Z_n(y)}, \quad A \subset \Theta \times X$$
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- (3) Robust to misspecification, robust statistics.
- (4) Calibration/violation of likelihood principle $P_n(A \mid y) = \frac{\int_A \exp(-\psi \ell_n(x,y)) d\pi(x)}{Z_n(y)}.$

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A SFT X is mixing if and only if there exists $N \ge 1$ with matrix

$$A_{uv} = \begin{cases} 1, & \text{if } \exists x \in \mathcal{X} \text{ with } x_{n-1} = u, x_n = v \\ 0 & \text{otherwise}, \end{cases}$$

and A^N contains all positive entries.

A Gibbs measure μ is given by a potential function $f : \mathcal{X} \to \mathbb{R}$ if there exists constants $\mathcal{P} \in \mathbb{R}$ and K > 0 such that for all $x \in \mathcal{X}$ and $m \ge 1$

$$K^{-1} \leq \frac{\mu(x_0^{m-1})}{\exp\left(-\mathcal{P}m + \sum_{k=1}^{m-1} f(S^k(x))\right)} \leq K.$$

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Under mild conditions unique, ergodic $\mu \in \mathcal{M}(\mathcal{X}, \mathcal{S})$.

 $\mathcal{F} = \{f_{\theta} : \theta \in \Theta\}$, with Θ compact is a regular family if for all $\theta \in \Theta$, $x \in \mathcal{X}$, and $m \ge 1$

$$K^{-1} \leq \frac{\mu_{\theta}(x_0^{m-1})}{\exp\left(-\mathcal{P}(f_{\theta})m + \sum_{k=1}^{m-1} f_{\theta}(\mathcal{S}^k(x))\right)} \leq K.$$

Hidden SFT models

Let \mathcal{X} be a mixing SFT, $\{f_{\theta} : \theta \in \Theta\} \subset C^{r}(\mathcal{X})$ a regular family of Hölder potential functions and $\{\mu_{\theta} : \theta \in \Theta\}$ the corresponding Gibbs measures with prior Π_{o} fully supported on Θ . Also let $\psi_{\theta}(u \mid x)$ be the observation process, with regularity.

Given the prior Π_o and marginal likelihood $p_{\theta}(u_0^{n-1})$ for $E \in \Theta$

$$\Pi_n(E \mid u_0^{n-1}) = \frac{\int_E p_\theta(u_0^{n-1}) d\Pi_o(\theta)}{\int_\Theta p_\theta(u_0^{n-1}) d\Pi_o(\theta)}$$

Posterior consistency

Theorem (McGoff-M-Nobel) Let $E \subset \Theta$ be an open neighborhood of $[\Theta^*]$. Then

$$\lim_{n} \Pi_{n}(\Theta \setminus E \mid U_{0}^{n-1}) = 0, \quad \mathbb{P}_{\theta^{*}}^{U} - a.s.$$

Joinings and couplings

Definition (Joining)

Let (X, A, μ, T) and (Y, B, ν, S) be two dynamical systems. A joining of T and S is a probability measure λ on $X \times Y$, with marginals μ and ν respectively, and invariant to the product map $T \times S$.

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Definition (Coupling)

A coupling of two random variable X and X' taking values in (E, \mathcal{E}) is any pair of random variables (Y, Y') taking values in $(E \times E, \mathcal{E} \times \mathcal{E})$ whose marginals have the same distribution as X and X', $X \stackrel{D}{=} Y$ and $X' \stackrel{D}{=} Y'$.

 $\mathcal{J}(\mu,\nu)$ is the set of all joinings of $(\mathcal{X}, \mathcal{S}, \mu)$ and $(\mathcal{Y}, \mathcal{T}, \nu)$.

Define $\mathcal{J}(S : \nu) = \bigcup_{\mu} \mathcal{J}(\mu, \nu)$, where the union is over all *S*-invariant Borel probability measures $\mu \in M(\mathcal{X}, S)$.

Variational formulation of $Z_n(y)$ – average cost

Recall ν is the measure for T and $\lambda \in \mathcal{J}(S : \nu)$

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Recall u is the measure for T and $\lambda \in \mathcal{J}(S: \nu)$

Define $\lambda_y \in M(\mathcal{X})$ (λ "projected" onto $d\nu_y$)

$$\lambda = \int_{\mathcal{Y}} \lambda_{\mathbf{y}} \otimes \delta_{\mathbf{y}} \, d\nu(\mathbf{y}).$$

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Limiting average cost

$$\lim_{n\to\infty}\frac{1}{n}\int_{\mathcal{X}}c_n(x,y)\,d\lambda_y(x)=\int c\,d\lambda_y(x)$$

Variational formulation of $Z_n(y)$ – entropy term

Given two Borel probability measures π and μ on \mathcal{X} and a finite measurable partition ξ of \mathcal{X} .

Denote $\mu \prec_{\xi} \pi$ as $\pi(C) = 0 \Rightarrow \mu(C) = 0$ for $C \in \xi$.

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Define

$$L(\mu \parallel \pi, \xi) = \begin{cases} \sum_{C \in \xi} \mu(C) \log \pi(C), & \text{if } \mu \prec_{\xi} \pi \\ -\infty, & \text{otherwise,} \end{cases}$$

with $0 \cdot \log 0 = 0$.

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In spirit consider all finite measurable partitions ξ

$$F(\mu,\pi) = \sup_{\xi} L(\mu \parallel \pi,\xi).$$

Theorem (McGoff-M.-Nobel)

Suppose a Glbbs prior, then for ν almost every y,

$$\lim_{n\to\infty}-\frac{1}{n}\log Z_n(y)=\inf_{\lambda\in\mathcal{J}(S:\nu)}\left\{\int c\,d\lambda+F(\lambda,\mu_\theta)\right\},$$

and the infimum in the above expression is attained.

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The above is the rate function in the large deviation sense.

Bayes as a variational problem

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A way to write Bayes rule

$$\pi(\theta \mid \mathbf{x}) = \arg\min_{\mu} \left\{ \int_{\theta} \ell(\theta, \mathbf{x}) d\mu(\theta) + d_{\mathsf{KL}}(\mu, \pi) \right\}$$

Convergence

Proposition (McGoff-M.-Nobel)

Suppose a Glbbs prior and consider the pressure

$$\begin{split} P(\mu_{\theta},\nu) &= \inf_{\lambda \in \mathcal{J}(S:\nu)} \left\{ \int c \, d\lambda + F(\lambda,\mu_{\theta}) \right\} \\ P(\theta:\nu) &= \inf_{\mu \in \mathcal{M}(\mathcal{X}_{\theta},\mathcal{S}_{\theta})} P(\mu_{\theta},\nu), \\ \theta_{*} &= \arg\min_{\theta \in \Theta} P(\theta:\nu). \end{split}$$

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For all $\varepsilon > 0$

$$P(d(S_{\theta_*}, T) < \varepsilon) \rightarrow 1$$
 a.s as $n \rightarrow \infty$.

Reframes posterior consistency as two-stage process: first find the limiting variational problem, and then analyze this problem to address consistency.

Provides general framework and suite of tools from the thermodynamic formalism for analyzing asymptotic behavior of Gibbs posteriors.

Statistics questions.

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Dynamics questions.

- ► How far can the thermodynamic formalism be pushed?
- Under what conditions is there a limiting variational characterization?

Questions

Statistics questions.

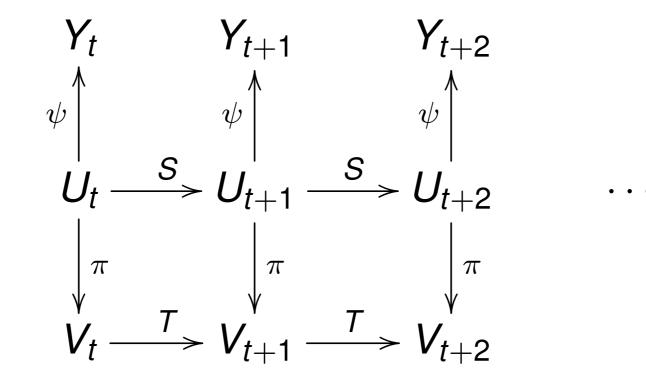
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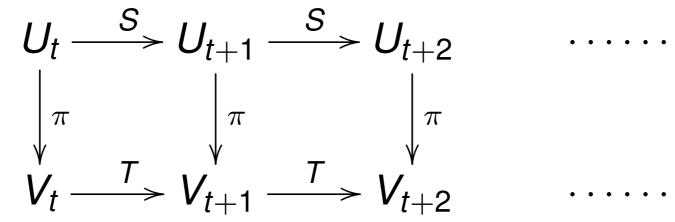
- ► How far can the thermodynamic formalism be pushed?
- Under what conditions is there a limiting variational characterization?
- Under what conditions is there a unique equilibrium joining?

Dimension reduction

Commutative diagrams

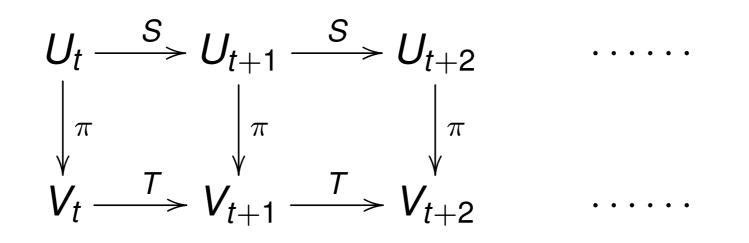


Topological conjugacy





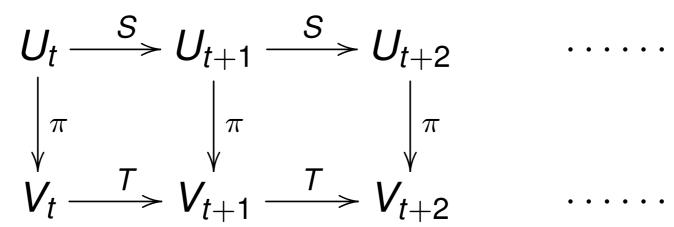
Topological conjugacy



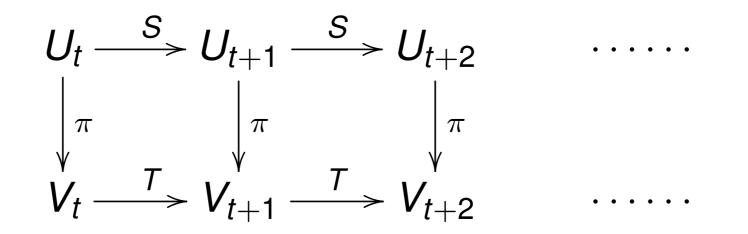
Topological conjugacy: Two functions $S : U \mapsto U$ and $T : V \mapsto V$ are *topologically conjugate* if there exists a homeomorphism $\pi : U \mapsto V$ such that

$$\pi \circ \boldsymbol{S} = \boldsymbol{T} \circ \pi.$$

Factors



Factors



Factors: Given two dynamical systems (U, \mathcal{U}, μ, T) and (V, \mathcal{V}, ν, S) with a map $\pi : U \mapsto V$ if

1. The map π is measurable.

2. For each $V \in \mathcal{V}$, $\mu(\pi^{-1}B) = \nu(B)$.

3. For μ -almost all $u \in U$, $\pi(Tx) = S(\pi x)$, Then (V, \mathcal{V}, ν, S) is a factor of (U, \mathcal{U}, μ, T) .

An objective

The factor suggests minimize either the difference in conditional probabilities

$$\min_{\pi^*, T} \mathsf{KL}(U_n \mid U_{n-1} \parallel \pi^*(V_n \mid V_{n-1}))$$

or the one step error

$$\min_{\pi^*, T} \mathbb{E} \| U_n - \pi^* (S(V_{n-1}) \|.$$

Partial factor model

Consider the the following linear regression setting for the dynamics

$$Y_i = B^T X_i + E_i, \quad E_i \stackrel{iid}{\sim} N(0, \Gamma)$$

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The statistical problem is to learn $A := \pi^*$ and B = T.

Typical joint distribution

$$egin{pmatrix} X \ Y \end{pmatrix} \sim \mathsf{N}(\mathbf{0}, \Sigma),$$

with

$$\Sigma = \begin{bmatrix} AA^T + \Gamma & BA^T \\ AB^T & \Psi + BB^T \end{bmatrix}$$

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The key idea

Instead model

$$\begin{pmatrix} X \\ f \\ Y \end{pmatrix} \sim \mathsf{N}(\mathbf{0}, \Sigma),$$

with

$$\Sigma = \begin{bmatrix} AA^T + \Gamma & A^T & BA^T \\ A & I_k & B^T \\ AB^T & B & \Psi + BB^T \end{bmatrix}$$

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We can apply some standard Bayesian models to infer A, B.

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- (5) Integration of ideas from statistical models of time series and dynamical systems theory.

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